Modern Computational Statistics

Lecture 12: Variational EM



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EM Recap

▶ EM algorithm finds the MLE for latent variable model

$$\mathcal{L}(\theta) = \log p(x|\theta) = \log \sum_{z} p(x, z|\theta)$$

▶ EM update formula

$$\theta^{(t+1)} = rg\max_{\theta} Q^{(t)}(\theta) = rg\max_{\theta} \mathbb{E}_{p(z|x,\theta^{(t)})} \log p(x, z|\theta)$$

- EM requires the posterior $p(z|x, \theta^{(t)})$ is known. What if $p(z|x, \theta^{(t)})$ is unknown?
 - If somehow we can sample from $p(z|x, \theta^{(t)})$, we can use Monte Carlo estimates, that is Monte Carlo EM.
 - ▶ However, the associated computation may be expansive.



Variational EM

▶ Recall EM maximizes the lower bound

$$\mathcal{F}(q, \theta^{(t)}) = \mathbb{E}_{q(z)} \log \frac{p(x, z|\theta)}{q(z)} \le \mathcal{L}(\theta), \quad \forall q(z)$$

- ► When the best $q(z) = p(z|x, \theta^{(t)})$ is not available, we can use approximate q(z) instead.
- ► A widely used approximation is the mean-field approximation

$$q(z) = \prod_{i=1}^{d} q_i(z_i)$$



Mean-Field Lower Bound

 \blacktriangleright In that case, the lower bound is

$$\mathcal{F}(q(z), \theta^{(t)}) = \int \prod_{i=1}^{d} q_i(z_i) \log \frac{p(x, z|\theta^{(t)})}{\prod_{i=1}^{d} q_i(z_i)} dz_1 dz_2 \dots dz_d$$
$$= \int \prod_{i=1}^{d} q_i(z_i) \log p(x, z|\theta^{(t)}) dz_1 dz_2 \dots dz_d$$
$$- \sum_{i=1}^{d} \int q_i(z_i) \log q_i(z_i) dz_i$$

Coordinate Ascent

$$q_i^{(t)}(z_i) \propto \exp\left(\mathbb{E}_{-q_i}\log p(x, z|\theta^{(t)})\right), i = 1, \dots, d$$



Mean-Field Variational EM

▶ E-step. Run coordinate ascent several times to obtain good mean-field approximation

$$q^{(t)}(z) = \prod_{i=1}^{d} q_i^{(t)}(z_i)$$

compute the expected complete data log-likelihood

$$Q^{(t)}(\theta) = \mathbb{E}_{q^{(t)}(z)} \log p(x, z|\theta)$$

• M-step. Update θ to maximize $Q^{(t)}(\theta)$

$$\theta^{(t+1)} = \underset{\theta}{\arg\max} Q^{(t)}(\theta)$$



Variational Bayesian EM

 Now let us consider Bayesian inference for latent variable models

$$p(z, \theta | x) \propto p(x, z | \theta) p(\theta)$$

▶ We can lower bound the marginal likelihood

$$\mathcal{L}(x) = \log p(x) = \log \int p(x, z|\theta) p(\theta) \, dz d\theta$$

= $\log \int q(z, \theta) \frac{p(x, z|\theta) p(\theta)}{q(z, \theta)} \, dz d\theta$
 $\geq \int q(z, \theta) \log \frac{p(x, z|\theta) p(\theta)}{q(z, \theta)} \, dz d\theta$
= $\mathcal{F}(q(z, \theta))$

• Maximizing this lower bound \mathcal{F} is equivalent to minimizing $D_{\mathrm{KL}}(q(z,\theta) \| p(z,\theta|x))$

Mean-Field Approximation

► Again, we consider a simple factorized approximation $q(z, \theta) = q_z(z)q_\theta(\theta)$

$$\mathcal{L}(x) \ge \int q_z(z)q_\theta(\theta)\log\frac{p(x,z|\theta)p(\theta)}{q_z(z)q_\theta(\theta)} \, dzd\theta$$
$$= \mathcal{F}(q_z(z),q_\theta(\theta))$$

► Maximizing this lower bound *F*, leads to **EM**-like iterative updates

$$q_{z}^{(t+1)}(z) \propto \exp\left(\mathbb{E}_{q_{\theta}^{(t)}(\theta)}\log p(x, z|\theta)\right)$$
$$q_{\theta}^{(t+1)}(\theta) \propto p(\theta) \cdot \exp\left(\mathbb{E}_{q_{z}^{(t+1)}(z)}\log p(x, z|\theta)\right)$$



Conjugate-Exponential Models

Let's focus on conjugate-exponential (CE) models, which satisfy Condition 1 The joint probability over variables is in the exponential family

$$p(x, z|\theta) = h(x, z) \exp \left(\phi(\theta) \cdot T(x, z) - A(\theta)\right)$$

Condition 2

The prior over parameters is conjugate to this joint probability

$$p(\theta|\eta,\nu) \propto \exp\left(\phi(\theta) \cdot \nu - \eta A(\theta)\right)$$

Conjugate priors are computationally convenient and have an intuitive interpretation:

- ▶ η : number of pseudo-observations
- ▶ ν : values of pseudo-observations



Conjugate-Exponential Models

Now suppose we have an iid data set $x = \{x_1, \ldots, x_n\}$ \blacktriangleright VB E-step.

$$q_z^{(t+1)}(z) \propto \exp\left(\mathbb{E}_{q_{\theta}^{(t)}(\theta)} \log p(x, z|\theta)\right)$$
$$\propto \prod_{i=1}^n h(x_i, z_i) \exp\left(\bar{\phi} \cdot T(x_i, z_i)\right)$$

where
$$\bar{\phi} = \mathbb{E}_{q_{\theta}^{(t)}}(\phi(\theta))$$

 \blacktriangleright VB M-step

$$q_{\theta}^{(t+1)}(\theta) \propto \exp\left(\phi(\theta) \cdot \left(\nu + \sum_{i=1}^{n} \overline{T}(x_i, z_i)\right) - (\eta + n)A(\theta)\right)$$

where $\overline{T}(x_i, z_i) = \mathbb{E}_{q_z^{(t+1)}}(T(x_i, z_i))$



EM for MAP v.s. Variational Bayesian EM

EM for MAP

- Goal: maximize $p(x, \theta)$
- **E-step**: compute

 $q_z^{(t+1)}(z) = p(z|x,\theta^{(t)})$

► M-step:

$$\theta^{(t+1)} = \operatorname*{arg\,max}_{\theta} Q^{(t)}(\theta)$$
$$Q^{(t)}(\theta) = \mathbb{E}_{q_z^{(t+1)}} \log p(x, z, \theta)$$

Variational Bayesian EM

- Goal: lower bound p(x)
- ► VB E-step: compute

$$q_z^{(t+1)}(z) = p(z|x,\bar{\phi})$$

► VB M-step:

 $\begin{aligned} q_{\theta}^{(t+1)}(\theta) &\propto \exp\left(Q^{(t)}(\theta)\right) \\ Q^{(t)}(\theta) &= \mathbb{E}_{q_{z}^{(t+1)}}\log p(x, z, \theta) \end{aligned}$



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- Reduces to the EM algorithm if $q_{\theta}(\theta) = \delta(\theta \theta^*)$.
- ► *F* increases monotonically, and incorporates the model complexity penalty.
- ► Analytical parameter distributions
- ▶ VB E-step has the same complexity as corresponding E step, and is almost identical except that it uses the expected natural parameters, $\bar{\phi}$.
- ▶ The lower bound given by VBEM can be used for model selection.

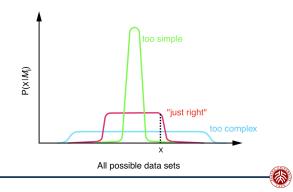


Bayesian Model Selection

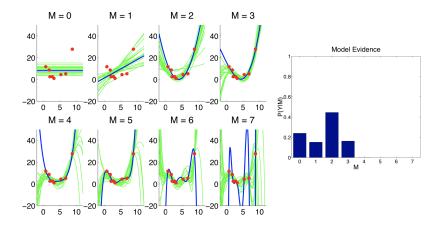
► In Bayesian model selection, we want to select the model class with the highest marginal likelihood (evidence)

$$p(x|m) = \int p(x|\theta, m) p(\theta|m) d\theta$$

▶ Occam's Razor



Bayesian Model Selection



Adapted from Zoubin Ghahramani



Bayesian Information Criterion (BIC):

$$\log p(x|m) \approx \log p(x|\hat{\theta}_{\text{MAP}}, m) - \frac{d}{2}\log n$$

► Annealed Importance Sampling (AIS):

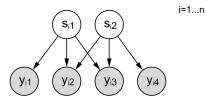
$$Z_k = \int p(x|\theta, m)^{\tau_k} p(\theta|m) d\theta, \quad 0 = \tau_0 < \dots < \tau_K = 1$$
$$\log p(x|\theta) = Z_K = \prod_{k=0}^{K-1} \frac{Z_{k+1}}{Z_k}$$

where Z_{k+1}/Z_k can be estimated via importance sampling.
▶ Variational Bayesian EM (VB): use VBEM lower bound estimate



Example: A Bipartite Structured Model

► A simple bipartite graphical model: **two** binary hidden variables, and **four** five-valued discrete observed variables

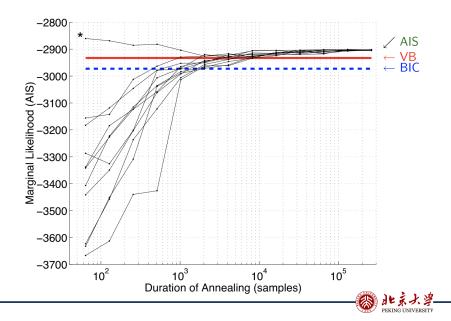


- Experiment: there are 136 distinct structures with 2 latent variables as potential parents of 4 conditionally independent observed variables
- Score each structure with 3 methods: BIC, VB and the gold standard AIS.



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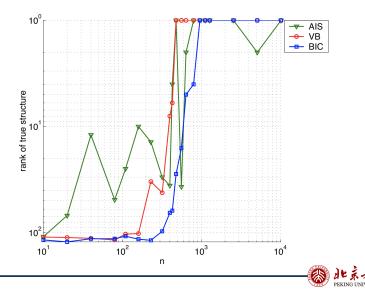
How Reliable is The AIS Gold Standard?



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Ranking The True Structure

VB score finds correct structure earlier, and more reliably



References

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- M. J. Beal and Z. Ghahramani, "The variational bayesian em algorithm for in- complete data: With application to scoring graphical model structures", Bayesian statistics, vol. 7, J. Bernardo, M. Bayarri, J. Berger, A. Dawid, D Heckerman, A. Smith, M West, et al., Eds., pp. 453–464, 2003.

